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HEPHY  
INSTITUTE OF  
HIGH ENERGY PHYSICS



European Research Council  
Established by the European Commission

using  at Belle II

invisible  $Z'$  search + tau LFU measurement

Paul Feichtinger

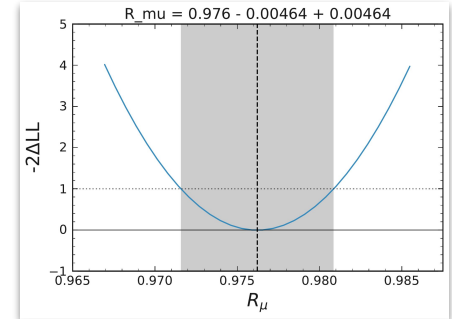
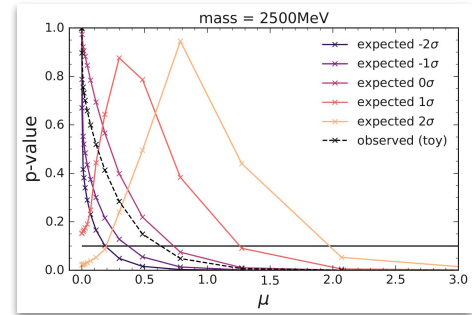
Belle II pyhf workshop

Bonn, 3rd March 2022

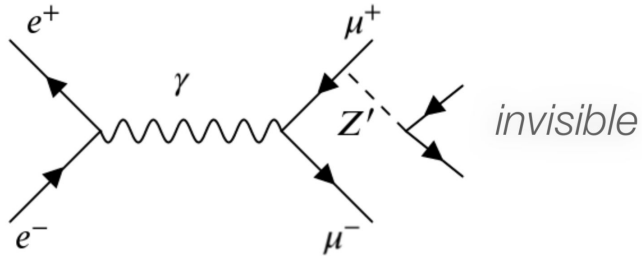
# Introduction

I used pyhf in two different analysis for

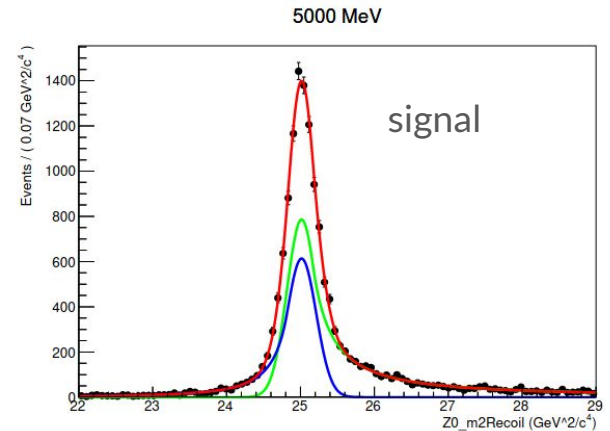
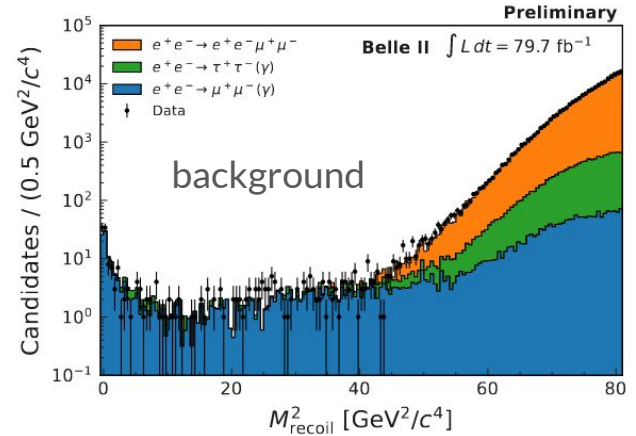
- search for an invisibly decaying  $Z'$  (BSM)
  - hypothesis testing
  
- test LFU in leptonic  $\tau$  decays (SM)
  - precision measurement

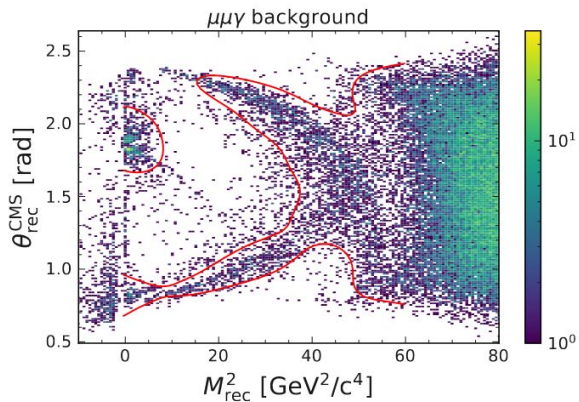


# Search for an invisibly decaying $Z'$

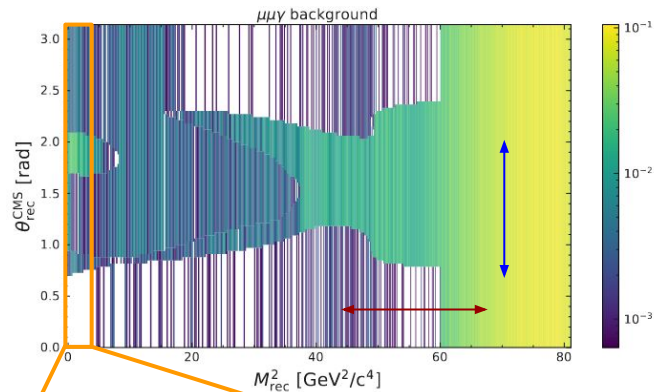


- search is bump hunt in recoil mass distribution
- perform a profile likelihood ratio test with pyhf
- challenges
  - many mass hypotheses to test ( $\sim 700$ )
  - low background environment

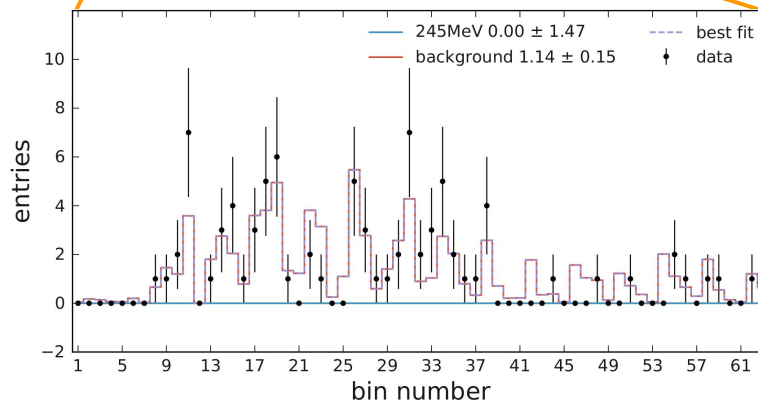




we derive a 2D template based on the **background distribution** and the **signal resolution**



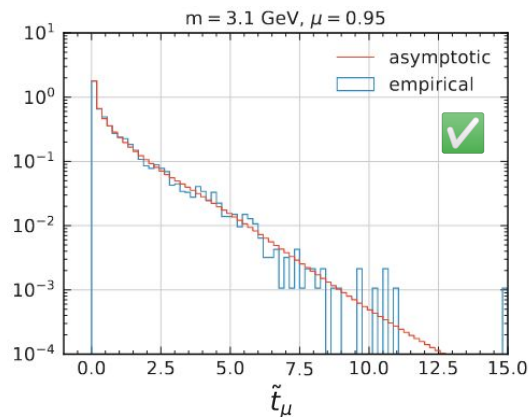
flattened bins for a 245MeV  $Z'$



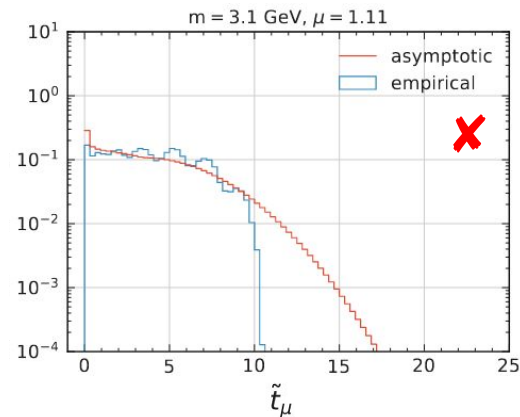
- many zero background bins in 2D template
  - raised error since  $\text{Pois}(n|\lambda=0)$  was not defined
  - removed 0 entry bins from the fit
  - should work since pyhf v0.7.0

- [asymptotic formulae](#) for distribution of test statistic only valid in the large sample limit
- easy to sample from the model to get empirical distribution of test statistic:  
`model.make_pdf(pars).sample()`
- background-like distributions were not accurately described by the asymptotics
  - used toys to derive expected limits

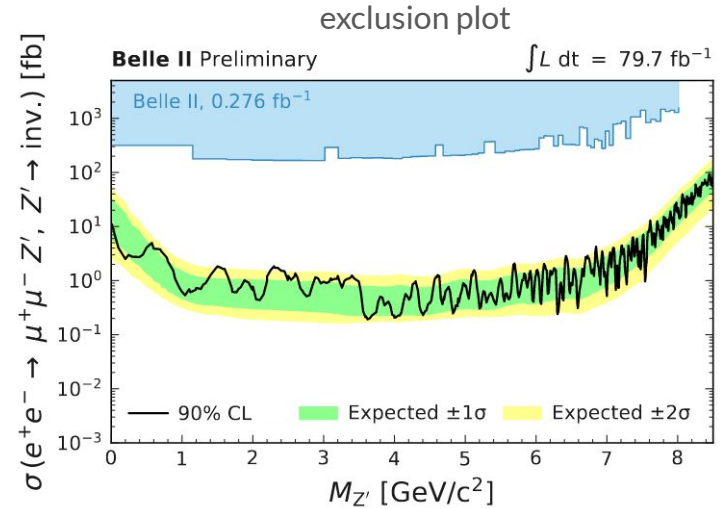
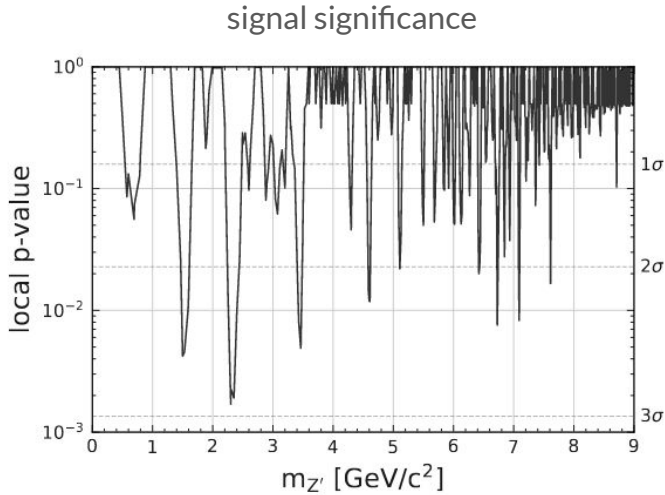
signal-like distribution (true POI = test POI)



background-like distribution (true POI = 0)



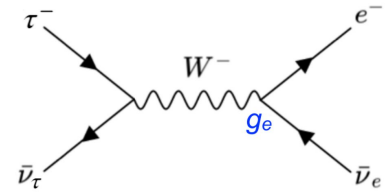
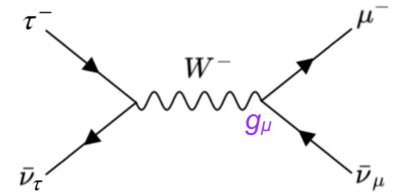
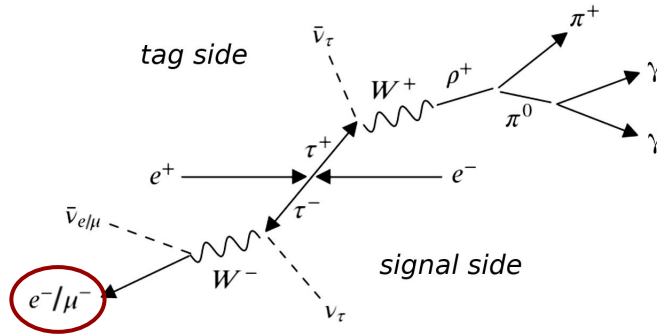
- compute expected + observed confidence intervals on cross section, for each mass hypothesis
- found that jax backend + scipy optimizer was very fast and robust



[arXiv:2212.03066](https://arxiv.org/abs/2212.03066)

# Testing LFU in leptonic $\tau$ decays

- want to measure the ratio  $R_\mu = \frac{\mathcal{B}(\tau^- \rightarrow \nu_\tau \mu^- \bar{\nu}_\mu(\gamma))}{\mathcal{B}(\tau^- \rightarrow \nu_\tau e^- \bar{\nu}_e(\gamma))}$
- close to one in SM if weak couplings  $g_\mu$  and  $g_e$  are equal



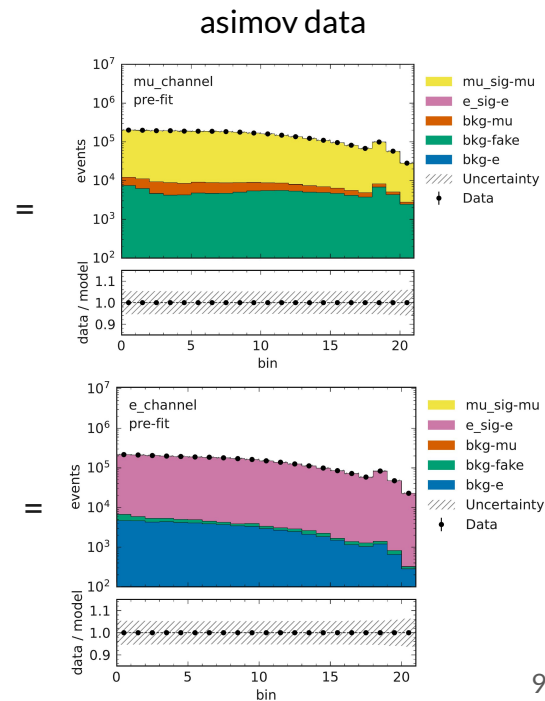
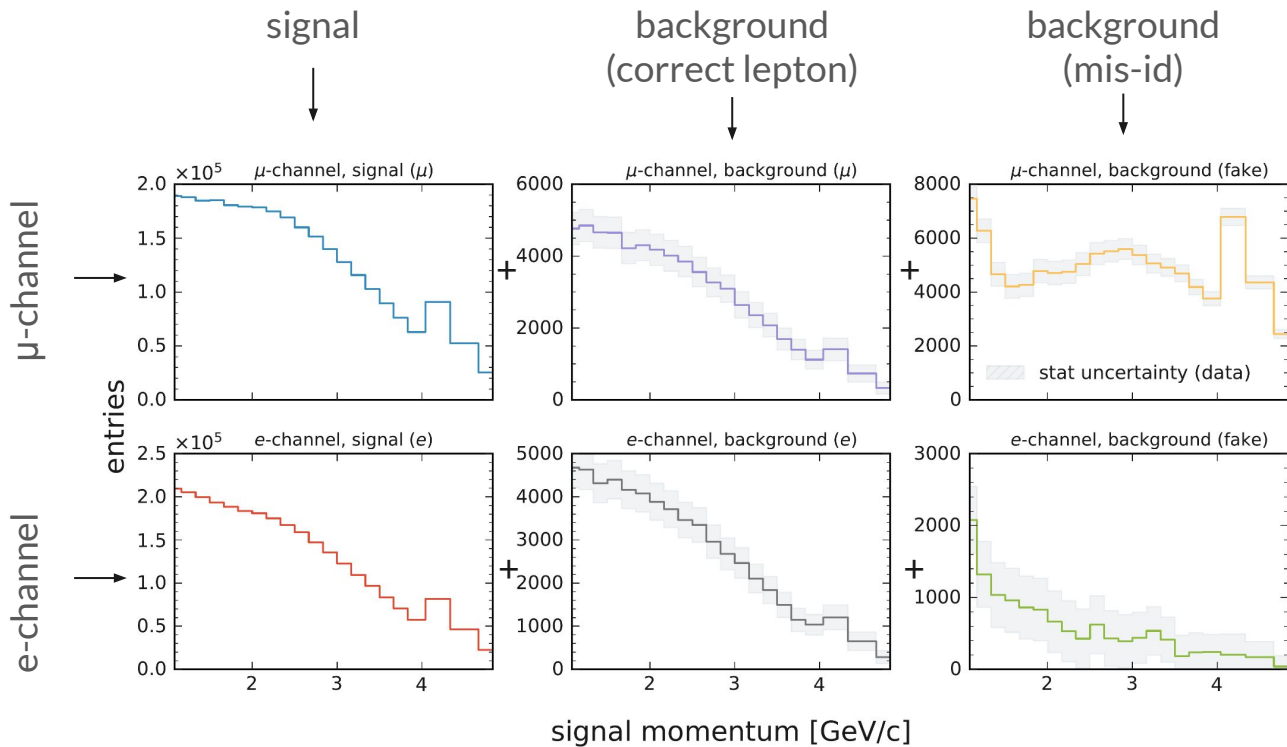
# Why do a template fit?

- we can achieve very high signal purity (~99%), so why do a template fit?
  - previous experiments only divided the background subtracted yields
  
- 1. electron / muon identification is leading systematic
  - lepton ID uncertainties heavily depend on lepton kinematics
  - by performing a fit in bins of lepton momentum we can reduce impact of regions with large uncertainty, without having to remove them
- 2. reduce uncertainty by providing as much information as possible
- 3. systematic uncertainties can be included with nuisance parameters directly in the fit
- 4. can publish full model with pyhf



# MC templates

separate events into 2 channels (based on lepton ID), each with 3 samples



# Systematic uncertainties

	uncertainty [%]	parameters	modifier
electron ID		7	histosys
control sample (signal embedding)		51*	shapesys
MC sample size		42	staterror
muon mis-ID		7	histosys
trigger		1+51*	normsys+shapesys
$\pi\pi^0$ modelling		113	shapesys
FSR (20% variation)		2	normsys+histosys
muon ID		7	histosys
electron mis-ID		7	histosys
normalisation		6	normsys
photon efficiency		1	histosys
$\pi^0$ efficiency		1	histosys
NN selection		1	histosys
misalignment		51*	shapesys
luminosity		1	lumi
ISR		2	normsys+histosys
decay in flight (5% variation)		1	histosys
$BF(\tau \rightarrow e\nu\bar{\nu})$		1	normsys
tracking efficiency		1	normsys
photon energy bias		1	histosys
track momentum		0	histosys
total systematic uncertainty			
statistical uncertainty			
total		253	

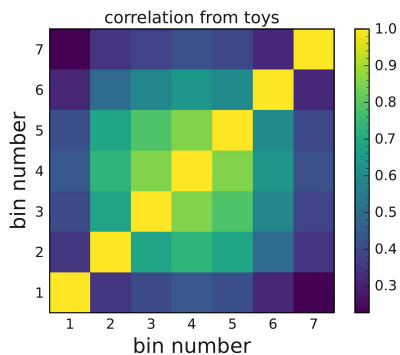
type of modifier depends on the correlation between template bins/samples

- **uncorrelated shape**: fully independent, one nuisance parameter per bin
- **correlated shape**: need to provide up/down variation, one nuisance parameter
  - can be shared for different samples
  - often cancels in the ratio
- **partial correlation** can be modeled with multiple correlated shape modifiers
  - using this for lepton ID

# Template variations from covariance

more details in backup

correlation for muon fake rate in 7 momentum bins

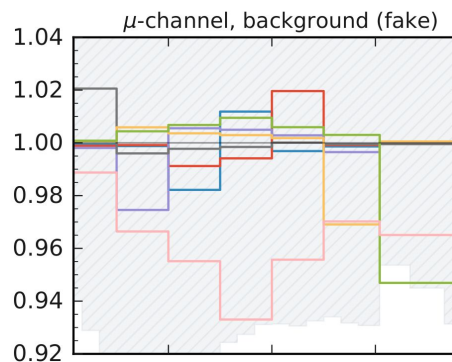


assumed correlation

eigendecomposition



$$C^{\text{sys}} = VUV^T$$



each colour shows a correlated shape variation (histosys)

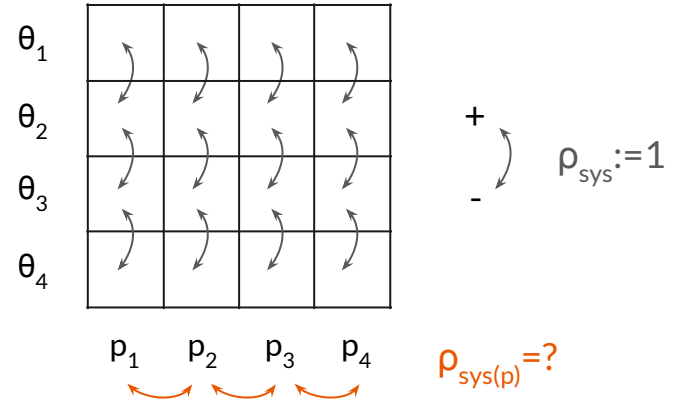
$$\Gamma = V\sqrt{U}$$

# Correlation of lepton ID systematics

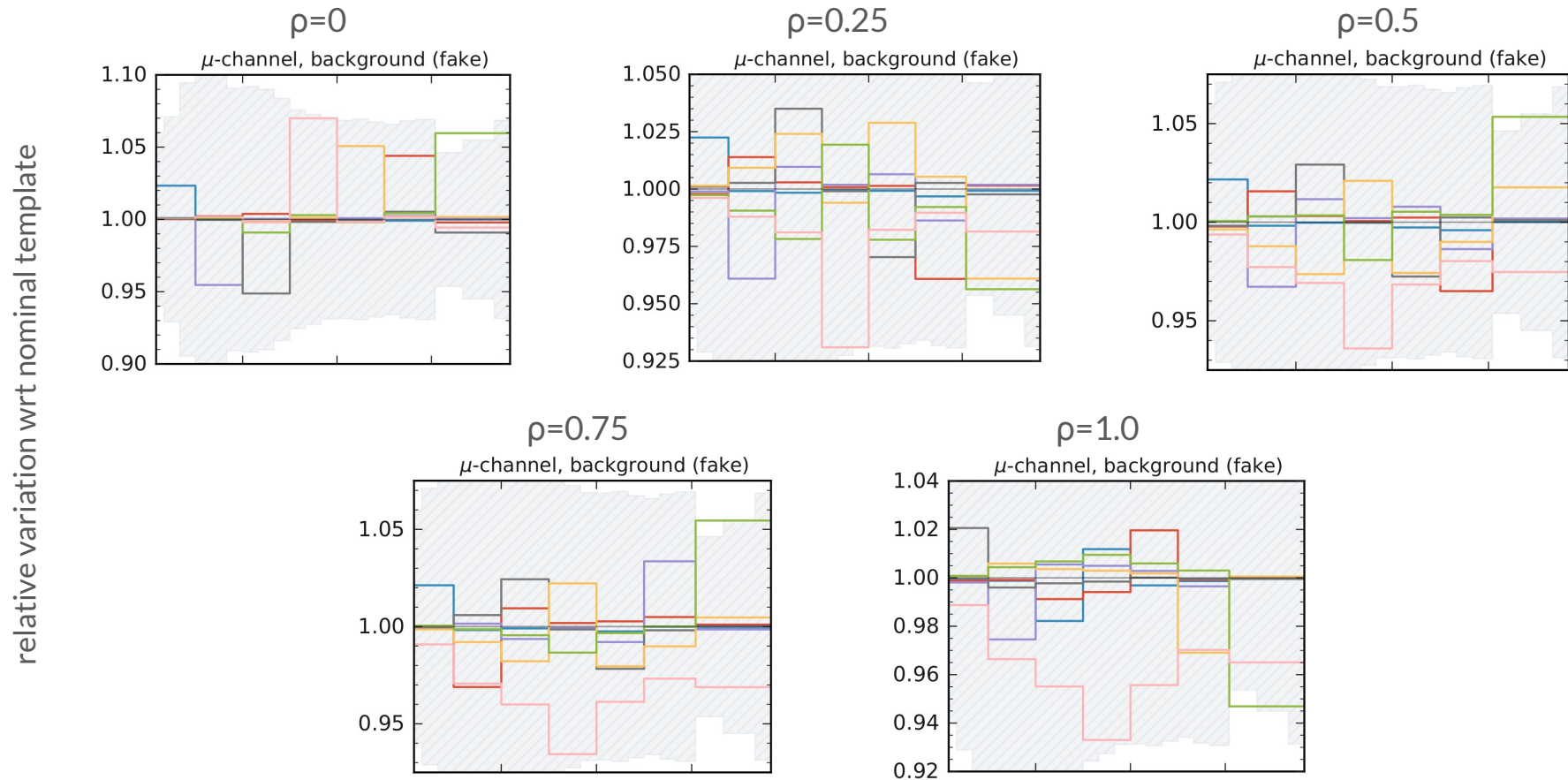
- LID uncertainties (stat + sys) are provided in bins of polar angle, momentum and charge
- statistical component of LID uncertainties is independent across correction bins
- correlation for the systematic component is currently unknown, need to make assumptions
  - want to be conservative and maintain a model that is able to describe the data

assuming correlation across...

- **charge**: fully correlated (conservative)
- **$\theta$** : fully correlated (conservative)
- **$p$** : not clear what is conservative, since fit is performed in momentum bins

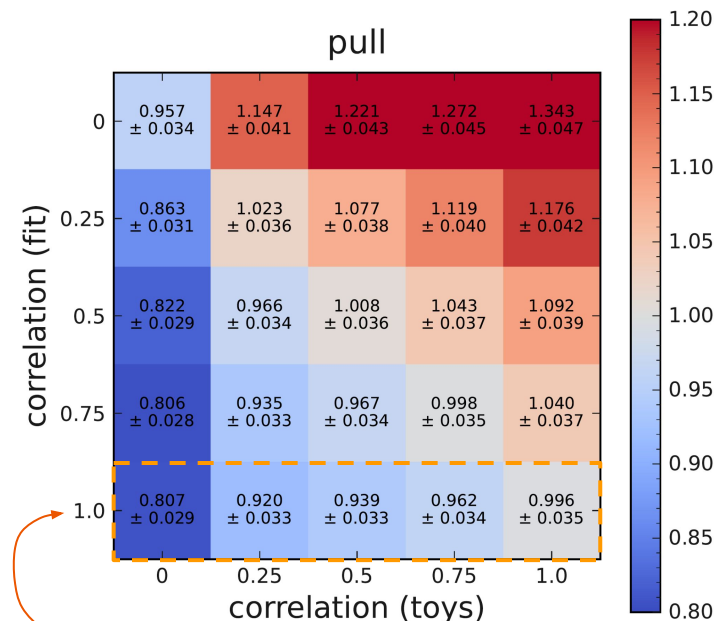
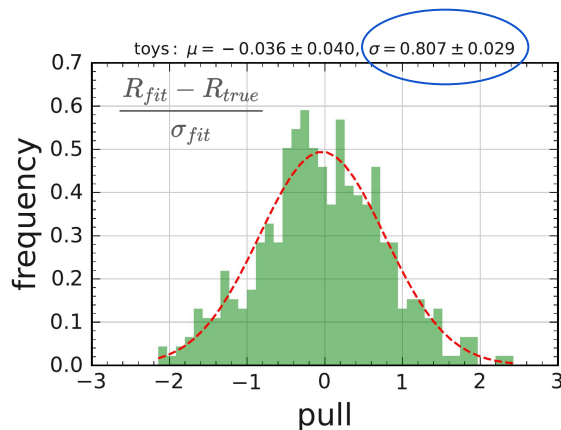


compute template variations for different correlations:  $\rho_{\text{sys}(p)} = \{0, 0.25, 0.5, 0.75, 1.0\}$



- want to identify the most conservative assumption
- we check the standard deviation of the pull distribution when fitting toys generated with one correlation with a model assuming a different correlation

pull when having  $\rho=0$  in toys  
 an  $\rho=1$  in model



$\rho_{\text{sys}(\rho)} = 1$  seems to be a good choice

independent variations probably 'absorbed' by other systematics (shapesys)

# Summary (why I use pyhf)

- easy to build a complicated statistical model out of well defined building blocks
- integrates well with different backends
- written in python
  - transparent: easy to read + modify code
  - interactive: can use jupyter
- limited to binned models, but flexible enough for many applications
  - can even model correlation, even though nuisance parameters are always independent

Thank you!

$$f(\mathbf{n}, \mathbf{a} | \boldsymbol{\eta}, \boldsymbol{\chi}) = \underbrace{\prod_{c \in \text{channels}} \prod_{b \in \text{bins}_c} \text{Pois}(n_{cb} | \nu_{cb}(\boldsymbol{\eta}, \boldsymbol{\chi}))}_{\text{Simultaneous measurement of multiple channels}} \underbrace{\prod_{\chi \in \boldsymbol{\chi}} c_{\chi}(a_{\chi} | \boldsymbol{\chi})}_{\text{constraint terms for "auxiliary measurements"}}$$

$$\nu_{cb}(\phi) = \sum_{s \in \text{samples}} \nu_{scb}(\boldsymbol{\eta}, \boldsymbol{\chi}) = \sum_{s \in \text{samples}} \underbrace{\left( \prod_{\kappa \in \boldsymbol{\kappa}} \kappa_{scb}(\boldsymbol{\eta}, \boldsymbol{\chi}) \right)}_{\text{multiplicative modifiers}} \left( \nu_{scb}^0(\boldsymbol{\eta}, \boldsymbol{\chi}) + \underbrace{\sum_{\Delta \in \boldsymbol{\Delta}} \Delta_{scb}(\boldsymbol{\eta}, \boldsymbol{\chi})}_{\text{additive modifiers}} \right)$$

Description	Modification	Constraint Term $c_{\chi}$	Input
Uncorrelated Shape	$\kappa_{scb}(\gamma_b) = \gamma_b$	$\prod_b \text{Pois}(r_b = \sigma_b^{-2}   \rho_b = \sigma_b^{-2} \gamma_b)$	$\sigma_b$
Correlated Shape	$\Delta_{scb}(\alpha) = f_p(\alpha   \Delta_{scb, \alpha=-1}, \Delta_{scb, \alpha=1})$	Gaus ( $a = 0   \alpha, \sigma = 1$ )	$\Delta_{scb, \alpha=\pm 1}$
Normalisation Unc.	$\kappa_{scb}(\alpha) = g_p(\alpha   \kappa_{scb, \alpha=-1}, \kappa_{scb, \alpha=1})$	Gaus ( $a = 0   \alpha, \sigma = 1$ )	$\kappa_{scb, \alpha=\pm 1}$
MC Stat. Uncertainty	$\kappa_{scb}(\gamma_b) = \gamma_b$	$\prod_b \text{Gaus}(a_{\gamma_b} = 1   \gamma_b, \delta_b)$	$\delta_b^2 = \sum_s \delta_{sb}^2$
Luminosity	$\kappa_{scb}(\lambda) = \lambda$	Gaus ( $l = \lambda_0   \lambda, \sigma_{\lambda}$ )	$\lambda_0, \sigma_{\lambda}$
Normalisation	$\kappa_{scb}(\mu_b) = \mu_b$		
Data-driven Shape	$\kappa_{scb}(\gamma_b) = \gamma_b$		



# Fit function

$$f(\vec{n} | R_\mu, \vec{\chi}) = \prod_{b \in \text{bins}} \text{Pois}(n_b^\mu | \nu_b^\mu(R_\mu, \vec{\chi})) \times \prod_{b \in \text{bins}} \text{Pois}(n_b^e | \nu_b^e(R_\mu, \vec{\chi})) \times \prod_{\chi \in \vec{\chi}} c_\chi(a_\chi | \chi)$$

data  $\nearrow$   $f(\vec{n} | R_\mu, \vec{\chi})$   
 parameter of interest  $\nearrow$   $R_\mu$   
 nuisance parameters  $\nearrow$   $\vec{\chi}$   
 $\mu$  channel  $\nearrow$   $\nu_b^\mu(R_\mu, \vec{\chi})$   
 $e$  channel  $\nearrow$   $\nu_b^e(R_\mu, \vec{\chi})$   
 constraints  $\nearrow$   $c_\chi(a_\chi | \chi)$

$$\nu_b^\mu(R_\mu, \vec{\chi}) = R_\mu \times \mathcal{B}_e \times \nu_b^{\mu, \text{sig}}(\vec{\chi}) + \nu_b^{\mu, \text{bkg}(\mu)}(\vec{\chi}) + \nu_b^{\mu, \text{bkg}(\text{fake})}(\vec{\chi})$$

$\nearrow = \mathcal{B}_\mu$

BF( $\tau \rightarrow e\nu\nu$ ) - multiplicative modifier

templates

$$\nu_b^e(R_\mu, \vec{\chi}) = \mathcal{B}_e \times \nu_b^{e, \text{sig}}(\vec{\chi}) + \nu_b^{e, \text{bkg}(e)}(\vec{\chi}) + \nu_b^{e, \text{bkg}(\text{fake})}(\vec{\chi})$$

# Template variations from covariance

- we can model any covariance matrix (for template bins) with multiple shape variations ([see PHYSTAT talk by S. Glazov](#))

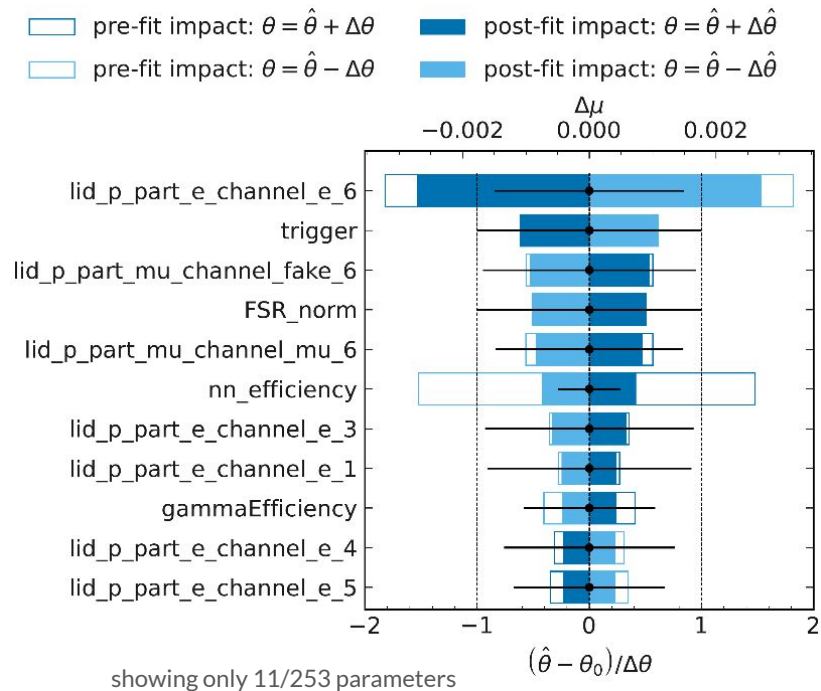
- relation between covariance matrix and template variations:  $C_{ik}^{\text{sys}} = \sum_{j=1}^{N_s} \Gamma_{ij} \Gamma_{kj}$
- $\Gamma_{ij}$  is a variation of bin i associated to systematic source j, in total  $N_s$  independent sources
- variations can be obtained from a covariance matrix via eigendecomposition

$$C^{\text{sys}} = VUV^T = (V\sqrt{U})(V\sqrt{U})^T = \Gamma\Gamma^T \implies \Gamma = V\sqrt{U}$$

- this results in as many shape variations as we have template bins
- can model any correlations also across samples

# Cabinetry

- [cabinetry](#) provides many convenience functions such as `cabinetry.fit.ranking()`
- pre-fit and post-fit impact on POI when varying nuisance parameters up/down one standard deviation
- can be used even if model is not created with `cabinetry`
- gives impact of each nuisance parameter individually, impact of a systematic source (e.g. lepton ID) can be obtained with toys



- test statistic  $\tilde{t}_\mu$  from [Asymptotic formulae for likelihood-based tests of new physics](#) (Cowan, Cranmer, Gross, Vitells)
- based on profile likelihood ratio,

$$\tilde{t}_\mu = -2 \ln \tilde{\lambda}(\mu) = \begin{cases} -2 \ln \left( L(\mu, \hat{\hat{\theta}}(\mu)) / L(0, \hat{\hat{\theta}}(0)) \right) & \hat{\mu} < 0 \\ -2 \ln \left( L(\mu, \hat{\hat{\theta}}(\mu)) / L(\hat{\mu}, \hat{\hat{\theta}}) \right) & \hat{\mu} \geq 0 \end{cases}$$

